

Lecture 23: 17 April, 2025

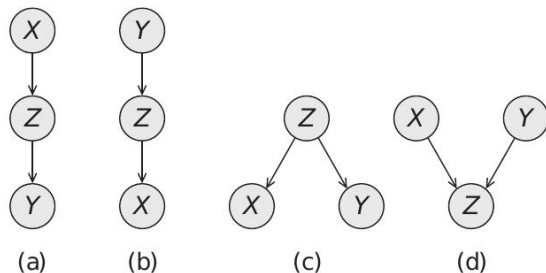
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Data Mining and Machine Learning
January–April 2025

D-Separation

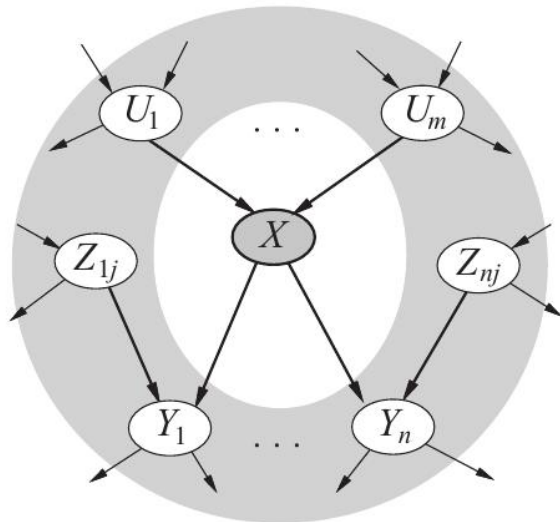
- Check if $X \perp Y \mid Z$
- Dependence should be blocked on every trail from X to Y
 - Each undirected path from X to Y is a sequence of basic trails
 - For (a), (b), (c), need Z present
 - For (d), need Z absent
 - In general, V-structure includes descendants of the bottom node



- x and y are **D-separated** given z if all trails are blocked
- Variation of **breadth first search (BFS)** to check if y is reachable from x through some trail
- Extends to sets — each $x \in X$ is D-separated from each $y \in Y$

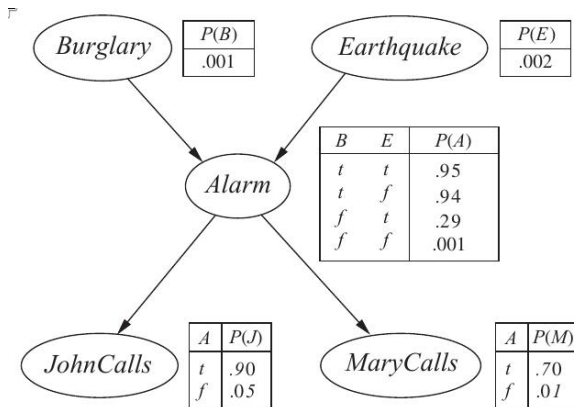
Markov blanket

- $MB(X)$ — Markov blanket of X
 - $Parents(X)$
 - $Children(X)$
 - $Parents\ of\ Children(X)$
- $X \perp \neg MB(X) \mid MB(X)$



Computing with probabilistic graphical models

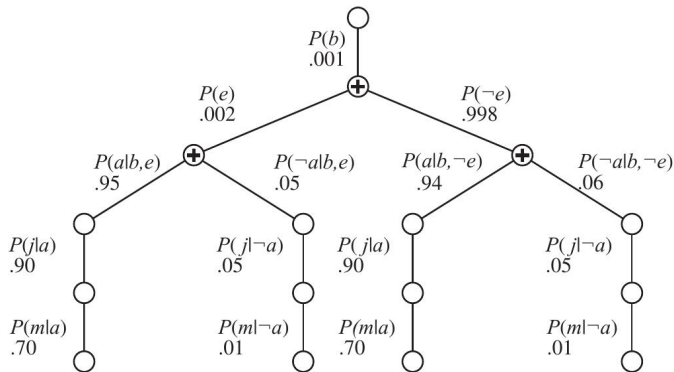
- John and Mary call Pearl. What is the probability that there has been a burglary?
- Want $P(b \mid m, j)$
- $\frac{P(b, m, j)}{P(m, j)}$
- Use chain rule to evaluate joint probabilities
- Reorder variables appropriately, topological order of graph



Computing with probabilistic graphical models

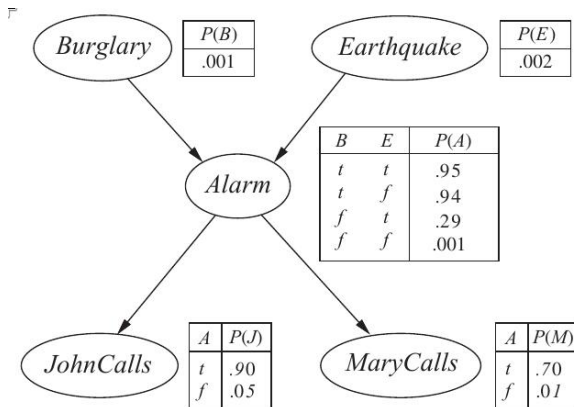
- $$P(m, j, b) = P(b) \sum_{e=0}^1 P(e) \sum_{a=0}^1 P(a | b, e) P(m | a) P(j | a)$$

- Construct the computation tree
- Use dynamic programming to avoid duplicated computations
- However, **exact inference** is NP-complete, in general
- Instead, **approximate inference** through sampling



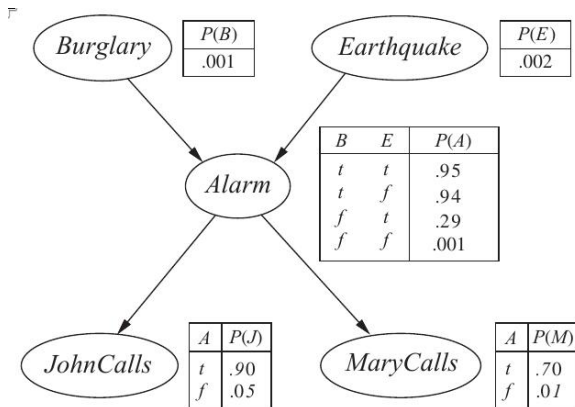
Approximate inference

- Generate random samples (b, e, a, m, j) , count to estimate probabilities
- Random samples should respect conditional probabilities
- Fix parents of x before generating x
- Generate in topological order
 - Generate b, e with probabilities $P(b)$ and $P(e)$
 - Generate a with probability $P(a \mid b, e)$
 - Generate j, m with probabilities $P(j \mid a)$, $P(m \mid a)$



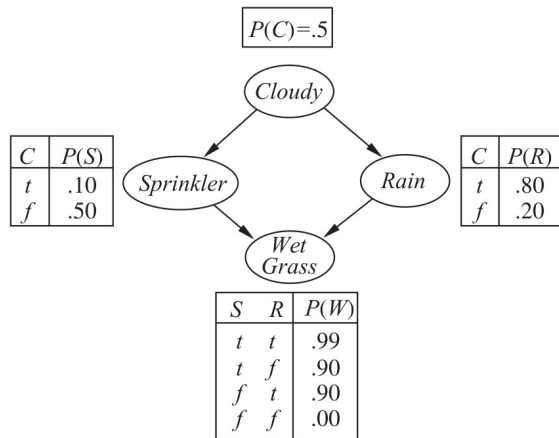
Approximate inference

- We are interested in $P(b \mid j, m)$
- Samples with $\neg j$ or $\neg m$ are useless
- Can we sample more efficiently?



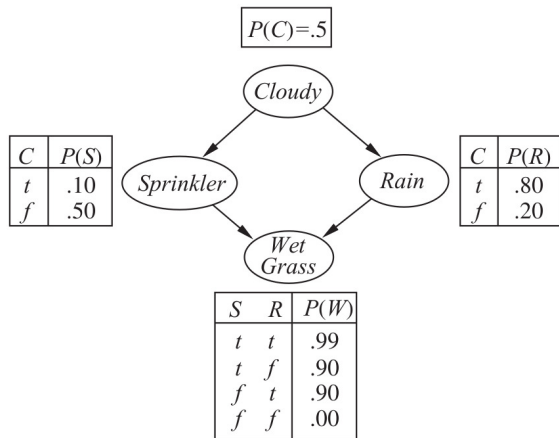
Rejection sampling

- $P(\text{Rain} \mid \text{Cloudy}, \text{Wet Grass})$
- Topological order
 - Generate *Cloudy*
 - Generate *Sprinkler*, *Rain*
 - Generate *Wet Grass*
- If we start with $\neg \text{Cloudy}$, sample is useless
- Immediately stop and reject this sample — rejection sampling
- General problem with low probability situation — many samples are rejected



Likelihood weighted sampling

- $P(\text{Rain} \mid \text{Cloudy}, \text{Wet Grass})$
- Fix **evidence** *Cloudy*, *Wet Grass* true
- Then generate the other variables
- Suppose we generate $c, \neg s, r, w$
- Compute likelihood of evidence:
 $0.5 \times 0.9 = 0.45$
- 0.45 is **likelihood weight** of sample
- Samples s_1, s_2, \dots, s_N with weights w_1, w_2, \dots, w_N
- $$P(r \mid c, w) = \frac{\sum_{s_i \text{ has rain}} w_i}{\sum_{1 \leq j \leq N} w_j}$$



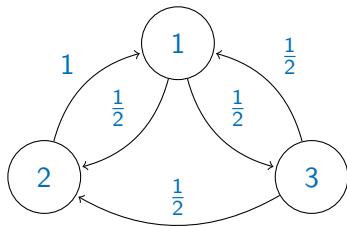
Gibbs sampling

- State of a Bayesian network is a valuation of variables (V_1, V_2, \dots, V_n)
- Move probabilistically from $s_j = (x_1, x_2, \dots, x_n)$ to $s_k = (y_1, y_2, \dots, y_n)$
- Allow such a move only when s_j, s_k differ at exactly one position
 - $s_j = (x_1, x_2, \dots, x_{i-1}, x_i, x_{i+1}, \dots, x_n)$
 - $s_k = (x_1, x_2, \dots, x_{i-1}, y_i, x_{i+1}, \dots, x_n)$
- Sampling algorithm
 - Current state is $s_j = (x_1, x_2, \dots, x_n)$
 - Choose i uniformly in $[1, n]$
 - Resample x_i given current values $(x_1, x_2, \dots, x_{i-1}, x_{i+1}, \dots, x_n)$
 - **Random walk** through state space — count number of visits to each state
- Need to compute $P[y_i \mid x_1, x_2, \dots, x_{i-1}, x_{i+1}, \dots, x_n]$
- Why does this work at all?

Approximate inference using Markov chains

Markov chains

- Finite set of states, with transition probabilities between states
- For us, a state will be an assignment of values to variables
- A three state Markov Chain



- Represent using a **transition matrix** — stochastic

$$A = \begin{bmatrix} 0 & \frac{1}{2} & \frac{1}{2} \\ 1 & 0 & 0 \\ \frac{1}{2} & \frac{1}{2} & 0 \end{bmatrix}$$

- $P[j]$ is probability of being in state j

- Start in state 1, so initially $P = \begin{bmatrix} 1 \\ 0 \\ 0 \end{bmatrix}$

Markov chains ...

- After one step:

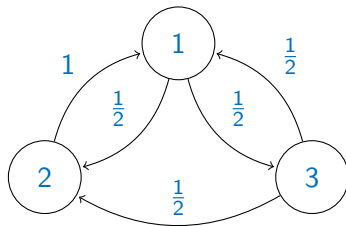
$$P^T A = \begin{bmatrix} 1 & 0 & 0 \end{bmatrix} \begin{bmatrix} 0 & \frac{1}{2} & \frac{1}{2} \\ 1 & 0 & 0 \\ \frac{1}{2} & \frac{1}{2} & 0 \end{bmatrix} = \begin{bmatrix} 0 & \frac{1}{2} & \frac{1}{2} \end{bmatrix}$$

- After second step:

$$\begin{bmatrix} 0 & \frac{1}{2} & \frac{1}{2} \end{bmatrix} \begin{bmatrix} 0 & \frac{1}{2} & \frac{1}{2} \\ 1 & 0 & 0 \\ \frac{1}{2} & \frac{1}{2} & 0 \end{bmatrix} = \begin{bmatrix} \frac{3}{4} & \frac{1}{4} & 0 \end{bmatrix}$$

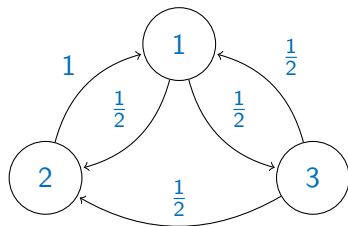
- After k steps, $P[j]$ is probability of being in state j
- Continuing our example,

$$\begin{bmatrix} \frac{3}{4} & \frac{1}{4} & 0 \end{bmatrix} \rightarrow \begin{bmatrix} \frac{1}{4} & \frac{3}{8} & \frac{3}{8} \end{bmatrix} \rightarrow \begin{bmatrix} \frac{9}{16} & \frac{5}{16} & \frac{1}{8} \end{bmatrix}$$



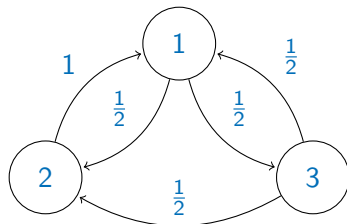
Ergodicity

- Is it the case that $P[j] > 0$ for all j continuously, after some point?
- Markov chain A is **ergodic** if there is some t_0 such that for every P , for all $t > t_0$, for every j , $(P^\top A^t)[j] > 0$.
 - No matter where we start, after $t > t_0$ steps, every state has a nonzero probability of being visited in step t
- Properties of ergodic Markov chains
 - There is a stationary distribution π , $\pi^\top A = \pi$
 - π is a **left eigenvector** of A
 - For *any* starting distribution P , $\lim_{t \rightarrow \infty} P^\top A^t = \pi$



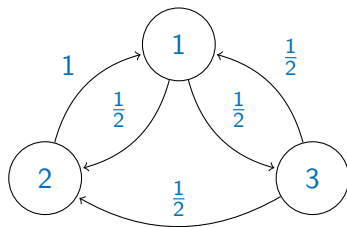
Ergodicity ...

- How can ergodicity fail?
 - Starting from i , we reach a set of states from which there is no path back to i
 - We have a cycle $i \rightarrow j \rightarrow k \rightarrow i \rightarrow j \rightarrow k \dots$, so we can only visit some states periodically
- Sufficient conditions for ergodicity
 - **Irreducibility**: When viewed as a directed graph, A is strongly connected
 - For all states i, j , there is a path from i to j and a path from j to i
 - **Aperiodicity**: For any pair of vertices i, j , the gcd of the lengths of all paths from i to j is 1
 - In particular, paths (loops) from i to i do not all have lengths that are multiples of some $k \geq 2$ — prevents bad cycles



Ergodicity ...

- Can efficiently approximate $\lim_{t \rightarrow \infty} P^\top A^t$ by repeated squaring: $P^\top A^2, P^\top A^4, P^\top A^8, \dots, P^\top A^{2^k}, \dots$
 - **Mixing time** — how fast this converges to π
- Stationary distribution represents fraction of visits to each state in a long enough execution
- Can we create a Markov chain from a Bayesian network so that the stationary distribution is meaningful?



Approximate inference using Markov chains

- Bayesian network has variables v_1, v_2, \dots, v_n
- Each assignment of values to the variables is a state
- Set up a Markov chain on these states
- Gibbs sampling — **random walk** through state space, count visits to each state
- Stationary distribution should assign to state s the probability $P(s)$ in the Bayesian network
- How to reverse engineer the transition probabilities to achieve this?

