Lecture 23: 11 April, 2023

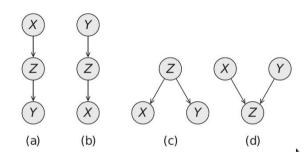
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Data Mining and Machine Learning January–April 2023

D-Separation

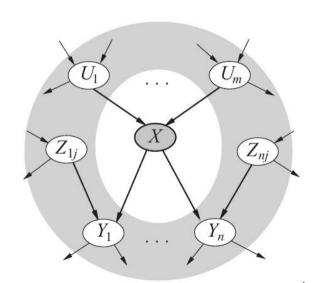
- Check if $X \perp Y \mid Z$
- Dependence should be blocked on every trail from X to Y
 - Each undirected path from X to Y is a sequence of basic trails
 - For (a), (b), (c), need Z present
 - For (d), need Z absent
 - In general, V-structure includes descendants of the bottom node



- x and y are D-separated given z if all trails are blocked
- Variation of breadth first search (BFS) to check if y is reachable from x through some trail
- Extends to sets each $x \in X$ is D-separated from each $y \in Y$

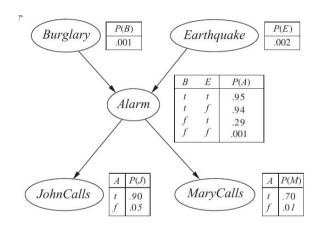
Markov blanket

- MB(X) Markov blanket of X
 - \blacksquare Parents(X)
 - Children(X)
 - Parents of Children(X)
- $\blacksquare X \perp \neg MB(X) \mid MB(X)$



Computing with probabilistic graphical models

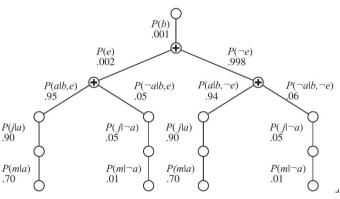
- John and Mary call Pearl. What is the probability that there has been a burglary?
- Want $P(b \mid m, j)$
- P(b, m, j)
- Use chain rule to evaluate joint probabilities
- Reorder variables appropriately, topological order of graph



Computing with probabilistic graphical models

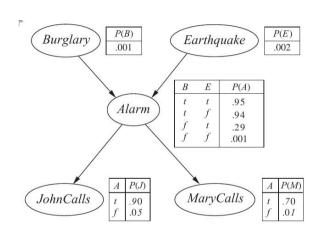
$$P(m,j,b) = P(b) \sum_{e=0}^{1} P(e) \sum_{a=0}^{1} P(a \mid b,e) P(m \mid a) P(j \mid a)$$

- Construct the computation tree
- Use dynamic programming to avoid duplicated computations
- However, exact inference is NP-complete, in general
- Instead, approximate inference through sampling



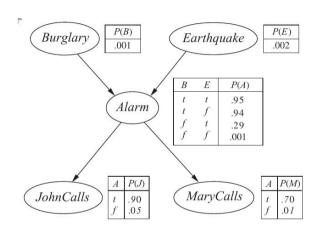
Approximate inference

- Generate random samples (b, e, a, m, j), count to estimate probabilities
- Random samples should respect conditional probabilities
- Fix parents of x before generating x
- Generate in topological order
 - Generate b, e with probabilities P(b) and P(e)
 - Generate *a* with probability $P(a \mid b, e)$
 - Generate j, m with probabilities $P(j \mid a)$, $P(m \mid a)$



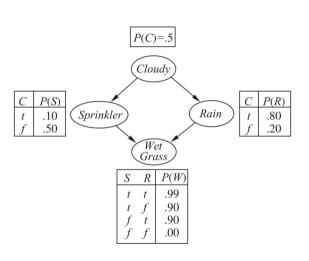
Approximate inference

- We are interested in P(b | j, m)
- Samples with $\neg j$ or $\neg m$ are useless
- Can we sample more efficiently?



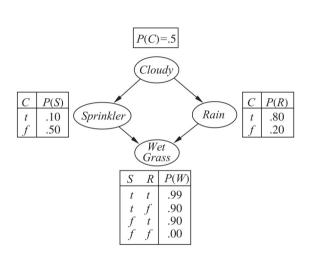
Rejection sampling

- *P*(*Rain* | *Cloudy*, *Wet Grass*)
- Topological order
 - Generate *Cloudy*
 - Generate Sprinkler, Rain
 - Generate Wet Grass
- If we start with $\neg Cloudy$, sample is useless
- Immediately stop and reject this sample — rejection sampling
- General problem with low probability situation — many samples are rejected



Likelihood weighted sampling

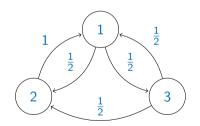
- *P*(*Rain* | *Cloudy*, *Wet Grass*)
- Fix evidence Cloudy, Wet Grass true
- Then generate the other variables
- Suppose we generate $c, \neg s, r, w$
- Compute likelihood of evidence: $0.5 \times 0.9 = 0.45$
- 0.45 is likelihood weight of sample
- Samples $s_1, s_2, ..., s_N$ with weights $w_1, w_2, ..., w_N$
- $P(r \mid c, w) = \frac{\sum_{s_i \text{ has rain } w_i}}{\sum_{1 \le i \le N} w_i}$



Approximate inference using Markov chains

Markov chains

- Finite set of states, with transition probabilities between states
- For us, a state will be an assignment of values to variables
- A three state Markov Chain



■ Represent using a transition matrix — stochastic

$$A = \left[\begin{array}{ccc} 0 & \frac{1}{2} & \frac{1}{2} \\ 1 & 0 & 0 \\ \frac{1}{2} & \frac{1}{2} & 0 \end{array} \right]$$

- P[j] is probability of being in state j
- Start in state 1, so initially $P = \begin{bmatrix} 1 \\ 0 \\ 0 \end{bmatrix}$

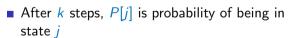
Markov chains . . .

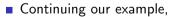
After one step:

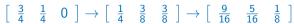
$$P^{\top}A = \begin{bmatrix} 1 & 0 & 0 \end{bmatrix} \begin{bmatrix} 0 & \frac{1}{2} & \frac{1}{2} \\ 1 & 0 & 0 \\ \frac{1}{2} & \frac{1}{2} & 0 \end{bmatrix} = \begin{bmatrix} 0 & \frac{1}{2} & \frac{1}{2} \end{bmatrix}$$

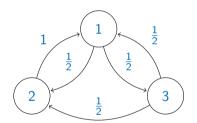
After second step:

$$\begin{bmatrix} 0 & \frac{1}{2} & \frac{1}{2} \end{bmatrix} \begin{bmatrix} 0 & \frac{1}{2} & \frac{1}{2} \\ 1 & 0 & 0 \\ \frac{1}{2} & \frac{1}{2} & 0 \end{bmatrix} = \begin{bmatrix} \frac{3}{4} & \frac{1}{4} & 0 \end{bmatrix}$$



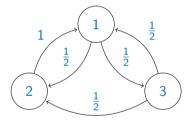






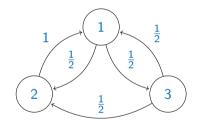
Ergodicity

- Is it the case that P[j] > 0 for all j continuously, after some point?
- Markov chain A is ergodic if there is some t_0 such that for every P, for all $t > t_0$, for every j, $(P^T A^t)[j] > 0$.
 - No matter where we start, after $t>t_0$ steps, every state has a nonzero probability of being visited in step t
- Properties of ergodic Markov chains
 - There is a stationary distribution π^* , $(\pi^*)^\top A = \pi^*$
 - \blacksquare π^* is a left eigenvector of A
 - For any starting distribution P, $\lim_{t\to\infty} P^{\top}A^t = \pi^*$



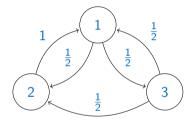
Ergodicity . . .

- How can ergodicity fail?
 - Starting from i, we reach a set of states from which there is no path back to i
 - We have a cycle $i \rightarrow j \rightarrow k \rightarrow i \rightarrow j \rightarrow k \cdots$, so we can only visit some states periodically
- Sufficient conditions for ergodicity
 - Irreducibility: When viewed as a directed graph, A is strongly connected
 - For all states *i*, *j*, there is a path from *i* to *j* and a path from *j* to *i*
 - Aperiodicity: For any pair of vertices *i*, *j*, the gcd of the lengths of all paths from *i* to *j* is 1
 - In particular, paths (loops) from i to i do not all have lengths that are multiples of some $k \ge 2$ prevents bad cycles



Ergodicity ...

- Can efficiently approximate $\lim_{t\to\infty} P^{\top}A^t$ by repeated squaring: $P^{\top}A^2$, $P^{\top}A^4$, $P^{\top}A^8$, ..., $P^{\top}A^{2^k}$, ...
 - Mixing time how fast this converges to π^*
- Stationary distribution represents fraction of visits to each state in a long enough execution
- Can we create a Markov chain from a Bayesian network so that the stationary distribution is meaningful?



Approximate inference using Markov chains

Bayesian network has variables

$$v_1, v_2, \ldots, v_n$$

- Each assignment of values to the variables is a state
- Set up a Markov chain based on these states
- Stationary distribution should assign to state s the probability P(s) in the Bayesian network
- How to reverse engineer the transition probabilities to achieve this?

