

Lecture 23: 28 April, 2022

Madhavan Mukund

<https://www.cmi.ac.in/~madhavan>

Data Mining and Machine Learning
January–May 2022

D-Separation

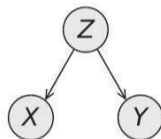
- Check if $X \perp Y \mid Z$
- Dependence should be blocked on every trail from X to Y
 - Each undirected path from X to Y is a sequence of basic trails
 - For (a), (b), (c), need Z present
 - For (d), need Z absent
 - In general, V-structure includes descendants of the bottom node



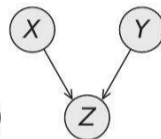
(a)



(b)



(c)

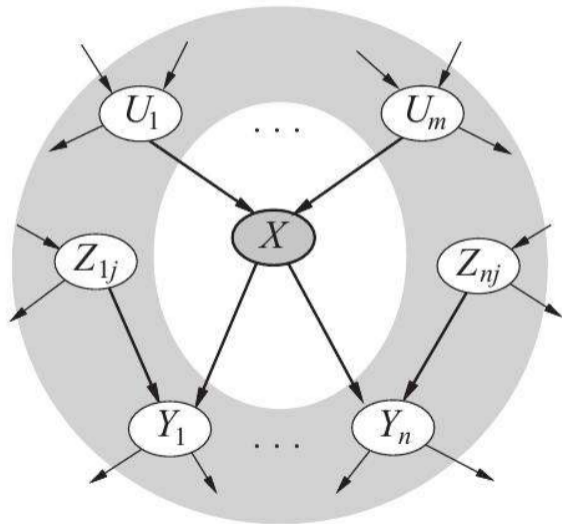


(d)

- x and y are **D-separated** given z if all trails are blocked
- Variation of **breadth first search (BFS)** to check if y is reachable from x through some trail
- Extends to sets — each $x \in X$ is D-separated from each $y \in Y$

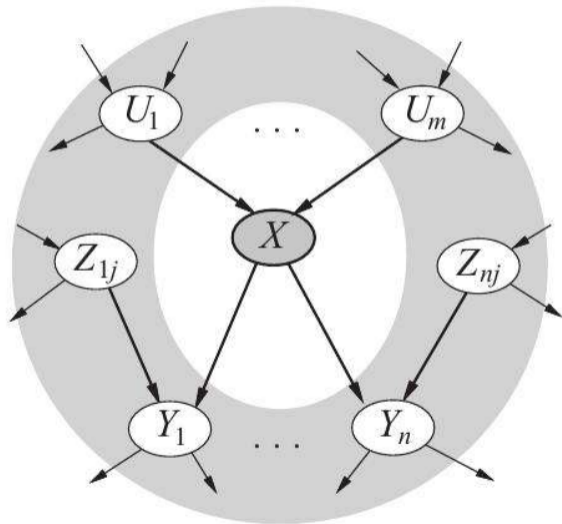
Markov blanket

- $MB(X)$ — Markov blanket of X



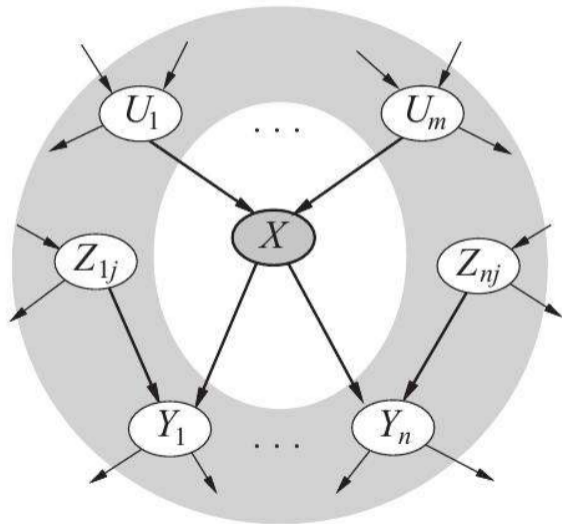
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- $MB(X)$ — Markov blanket of X
 - $Parents(X)$



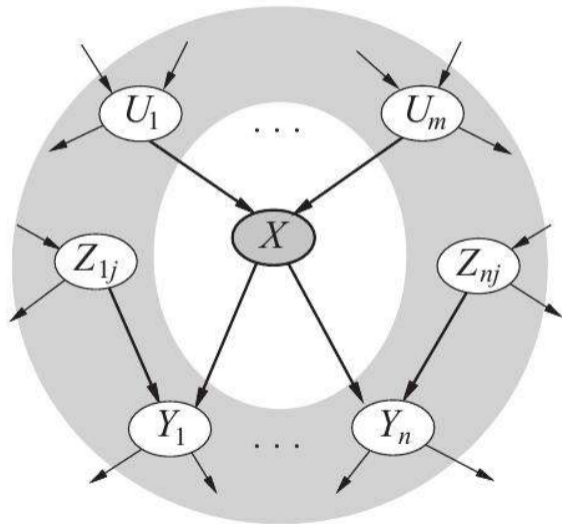
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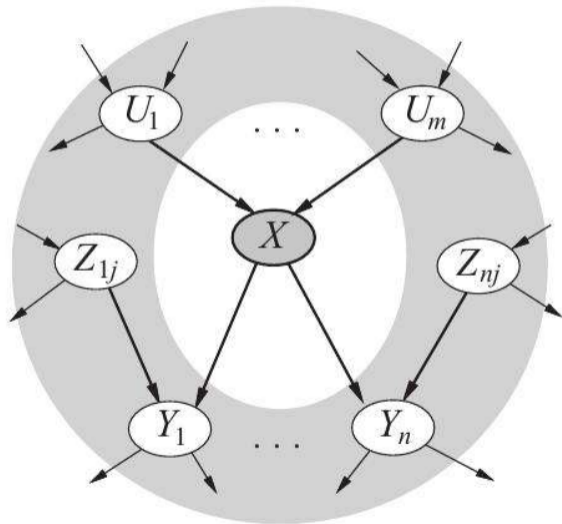
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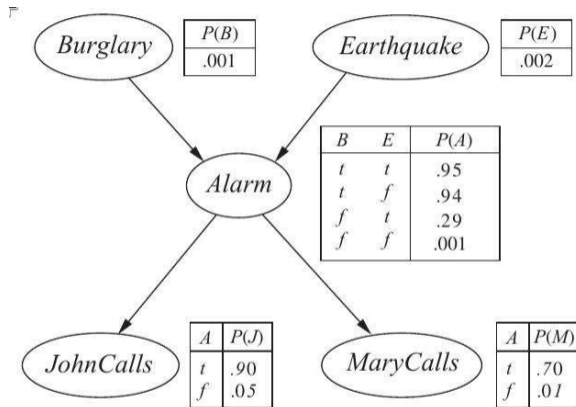
Markov blanket

- $MB(X)$ — Markov blanket of X
 - $Parents(X)$
 - $Children(X)$
 - $Parents\ of\ Children(X)$
- $X \perp \neg MB(X) \mid MB(X)$



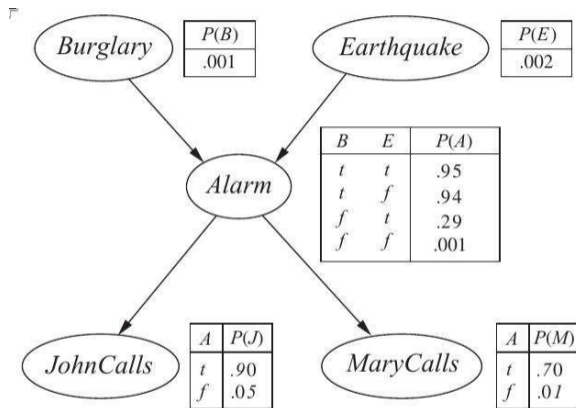
Computing with probabilistic graphical models

- John and Mary call Pearl. What is the probability that there has been a burglary?



Computing with probabilistic graphical models

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- Want $P(b | m, j)$

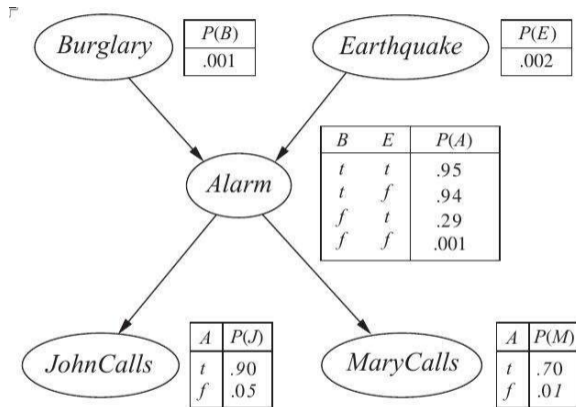


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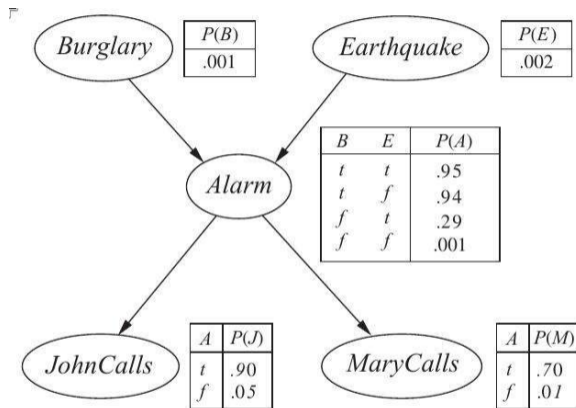
- Want $P(b \mid m, j)$

- $$\frac{P(b, m, j)}{P(m, j)}$$



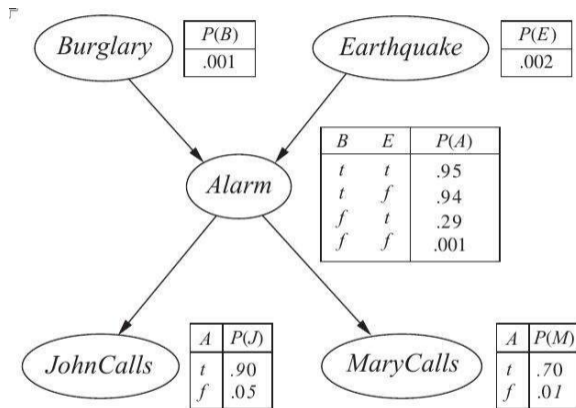
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- $\frac{P(b, m, j)}{P(m, j)}$
- Use chain rule to evaluate joint probabilities



Computing with probabilistic graphical models

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- Want $P(b \mid m, j)$
- $\frac{P(b, m, j)}{P(m, j)}$
- Use chain rule to evaluate joint probabilities
- Reorder variables appropriately, topological order of graph

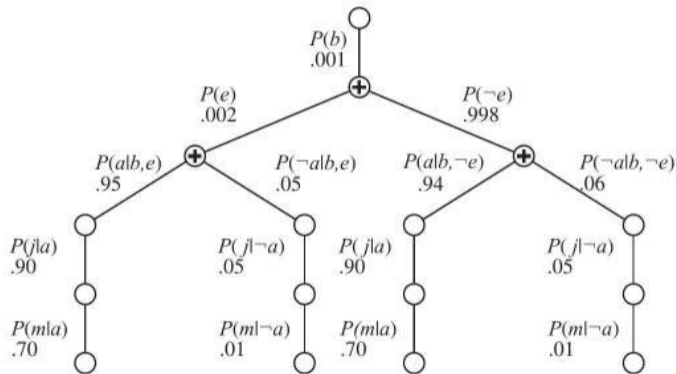


Computing with probabilistic graphical models

- $$P(m, j, b) = P(b) \sum_{e=0}^1 P(e) \sum_{a=0}^1 P(a | b, e) P(m | a) P(j | a)$$

Computing with probabilistic graphical models

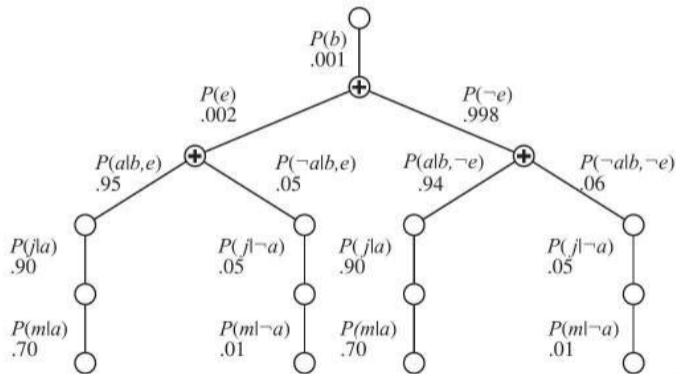
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- Construct the computation tree



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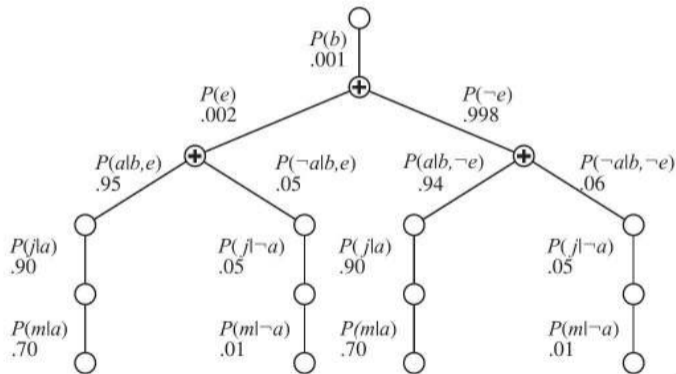
- Construct the computation tree
- Use dynamic programming to avoid duplicated computations



Computing with probabilistic graphical models

- $$P(m, j, b) = P(b) \sum_{e=0}^1 P(e) \sum_{a=0}^1 P(a | b, e) P(m | a) P(j | a)$$

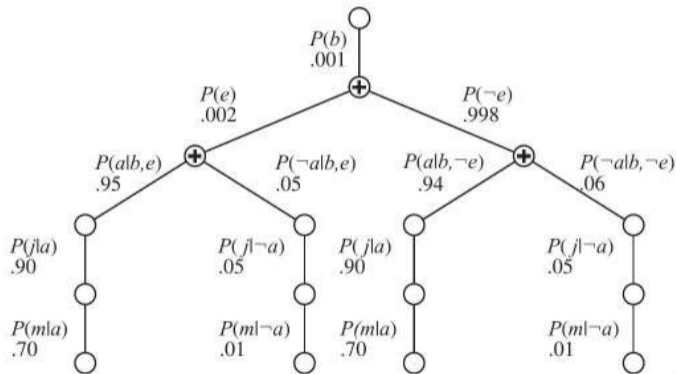
- Construct the computation tree
- Use dynamic programming to avoid duplicated computations
- However, **exact inference** is NP-complete, in general



Computing with probabilistic graphical models

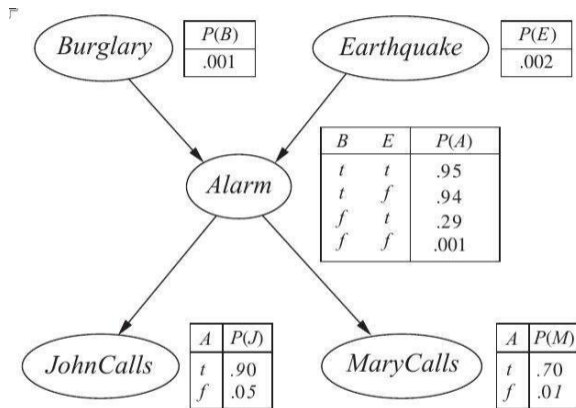
- $$P(m, j, b) = P(b) \sum_{e=0}^1 P(e) \sum_{a=0}^1 P(a | b, e) P(m | a) P(j | a)$$

- Construct the computation tree
- Use dynamic programming to avoid duplicated computations
- However, **exact inference** is NP-complete, in general
- Instead, **approximate inference** through sampling



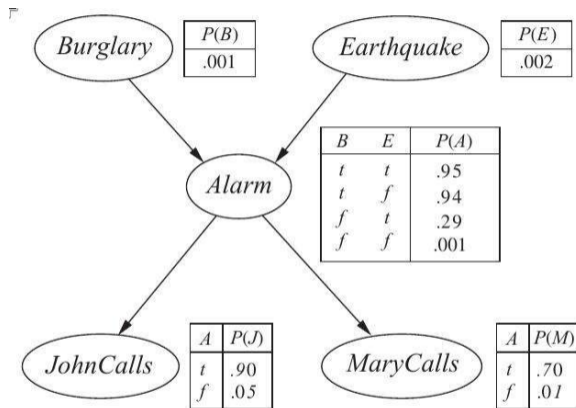
Approximate inference

- Generate random samples (b, e, a, m, j) , count to estimate probabilities



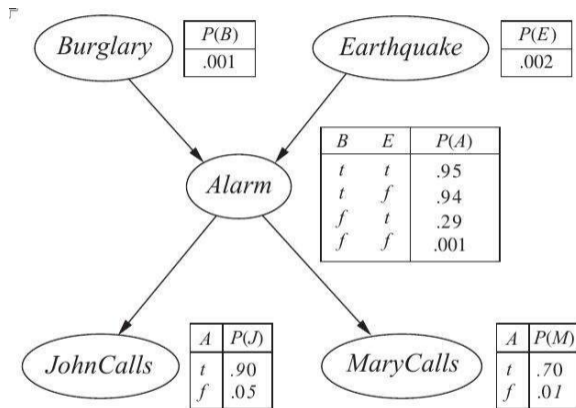
Approximate inference

- Generate random samples (b, e, a, m, j) , count to estimate probabilities
- Random samples should respect conditional probabilities



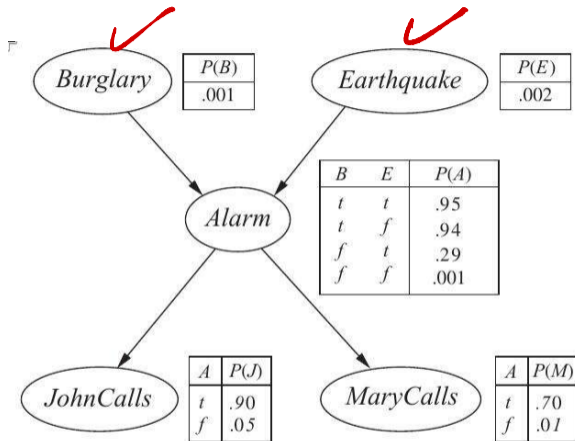
Approximate inference

- Generate random samples (b, e, a, m, j) , count to estimate probabilities
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- Fix $MB(x)$ before generating x



Approximate inference

- Generate random samples (b, e, a, m, j) , count to estimate probabilities
- Random samples should respect conditional probabilities
- ~~Fix $MB(x)$ before generating x~~
- Generate in topological order
 - Generate b, e with probabilities $P(b)$ and $P(e)$
 - Generate a with probability $P(a | b, e)$
 - Generate j, m with probabilities $P(j | a)$, $P(m | a)$

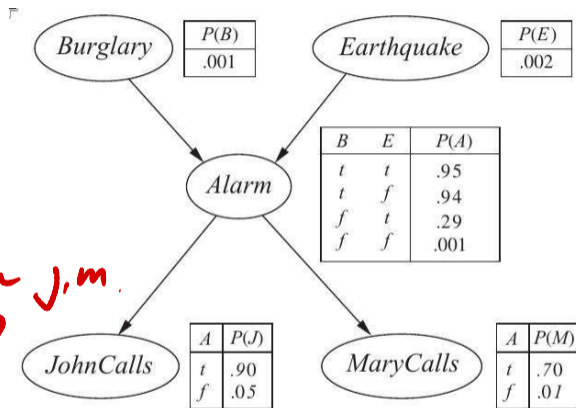


Approximate inference

- We are interested in $P(b | j, m)$

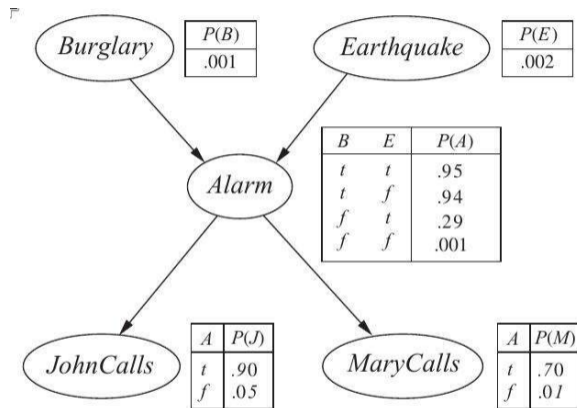
$$\frac{P(b, j, m)}{P(j, m)}$$

Among all samples with j, m .
how many have b ?



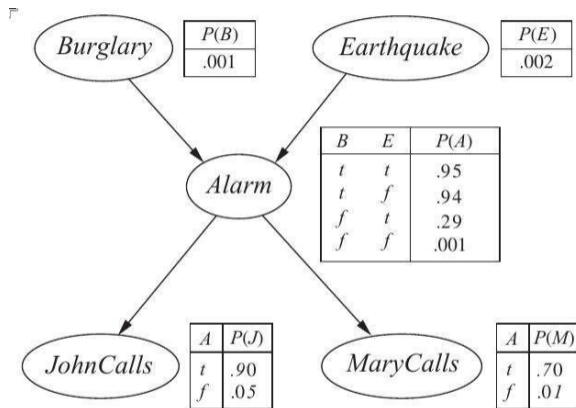
Approximate inference

- We are interested in $P(b \mid j, m)$
- Samples with $\neg j$ or $\neg m$ are useless



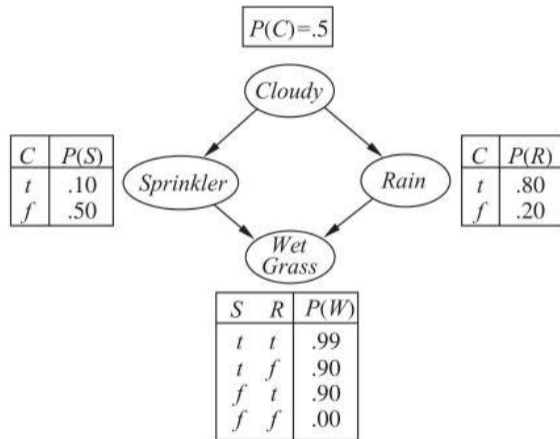
Approximate inference

- We are interested in $P(b | j, m)$
- Samples with $\neg j$ or $\neg m$ are useless
- Can we sample more efficiently?



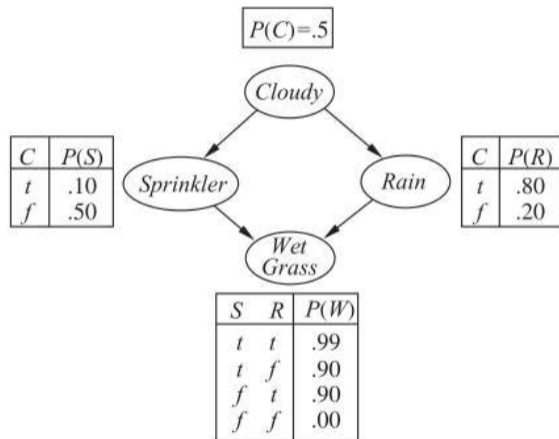
Rejection sampling

- $P(\text{Rain} \mid \text{Cloudy}, \text{Wet Grass})$



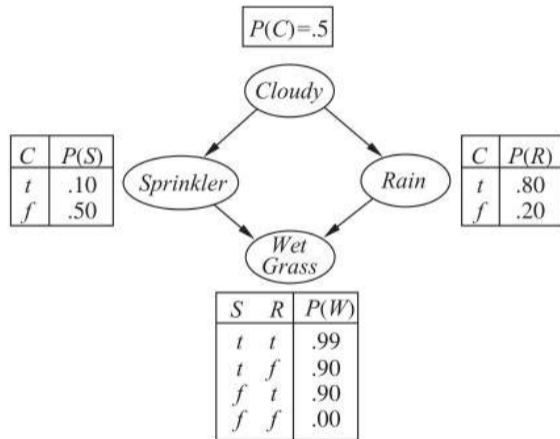
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- $P(\text{Rain} \mid \text{Cloudy}, \text{Wet Grass})$
- Topological order
 - Generate *Cloudy*
 - Generate *Sprinkler*, *Rain*
 - Generate *Wet Grass*



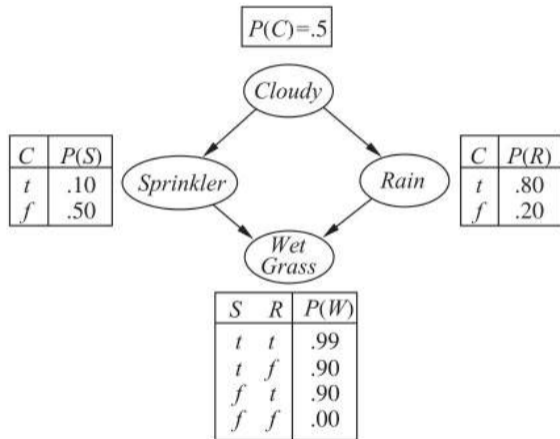
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- If we start with $\neg \text{Cloudy}$, sample is useless



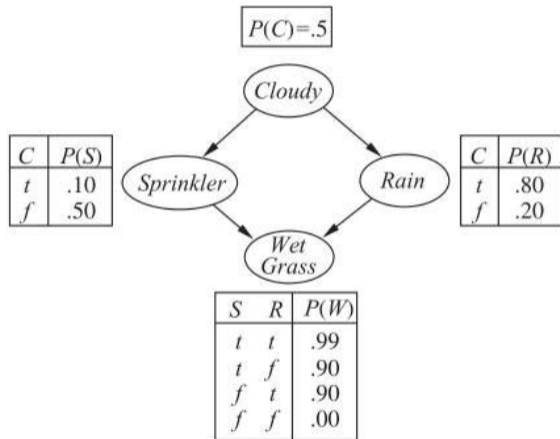
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- Immediately stop and reject this sample — **rejection sampling**



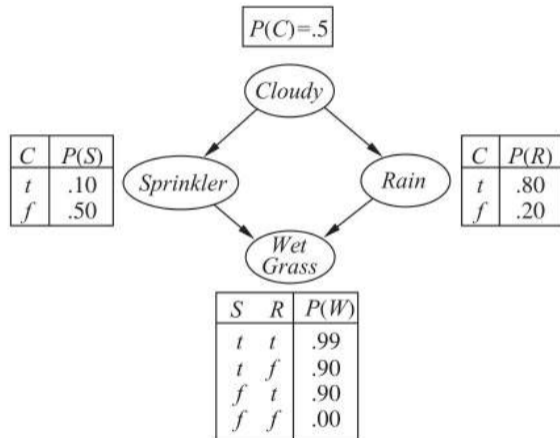
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- Immediately stop and reject this sample — **rejection sampling**
- General problem with low probability situation — lots of samples



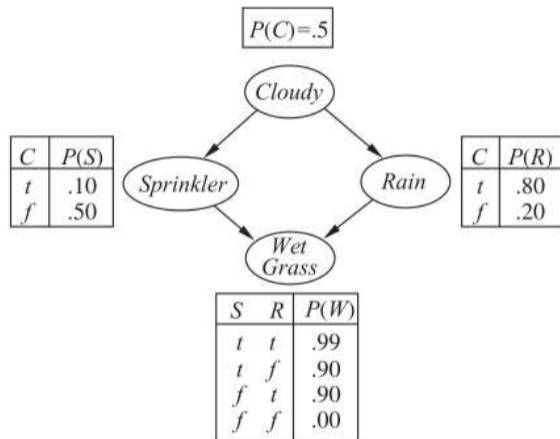
Likelihood weighted sampling

- $P(\text{Rain} \mid \text{Cloudy}, \text{Wet Grass})$



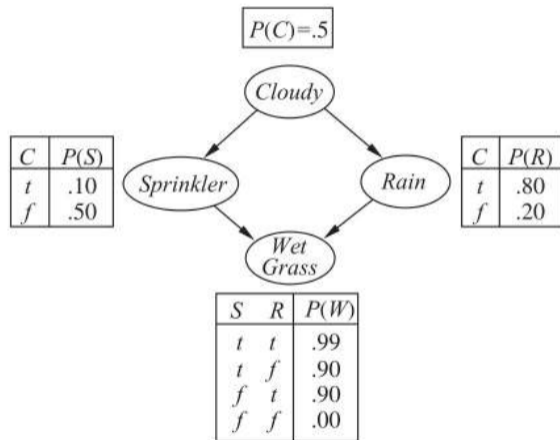
Likelihood weighted sampling

- $P(\text{Rain} \mid \text{Cloudy}, \text{Wet Grass})$
- Fix **evidence** *Cloudy, Wet Grass* true



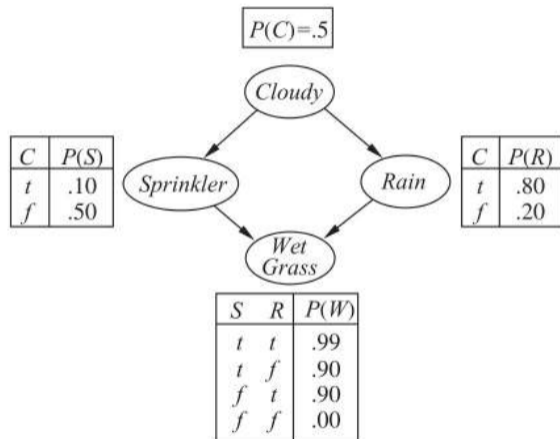
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- $P(\text{Rain} \mid \text{Cloudy}, \text{Wet Grass})$
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- Then generate the other variables



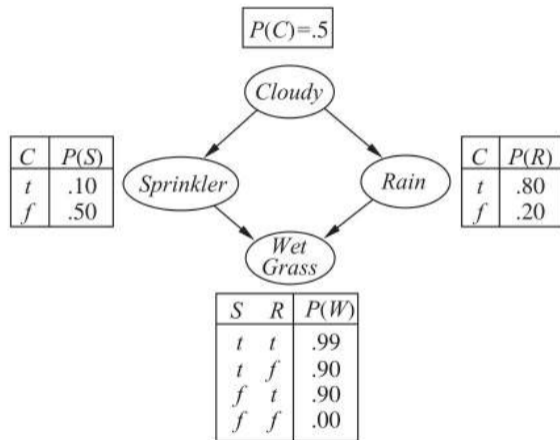
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- Suppose we generate $c, \neg s, r, w$



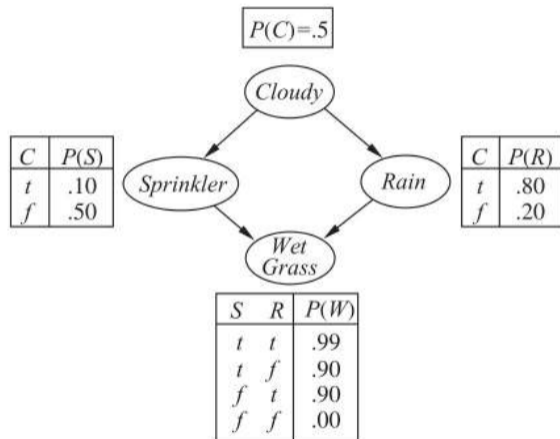
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- Compute likelihood of evidence:
 $0.5 \times 0.9 = 0.45$



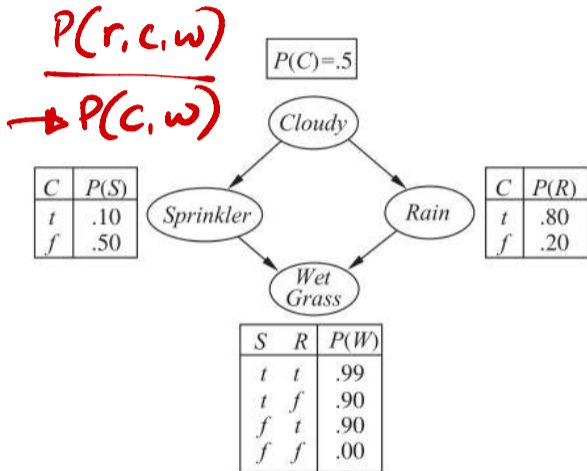
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- 0.45 is **likelihood weight** of sample



Likelihood weighted sampling

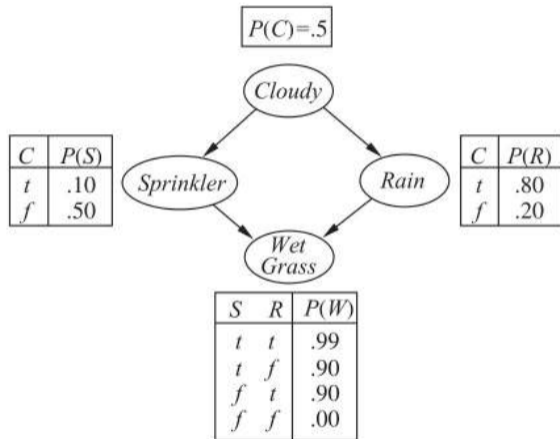
- $P(\text{Rain} \mid \text{Cloudy}, \text{Wet Grass})$
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- Samples s_1, s_2, \dots, s_N with weights
 w_1, w_2, \dots, w_N



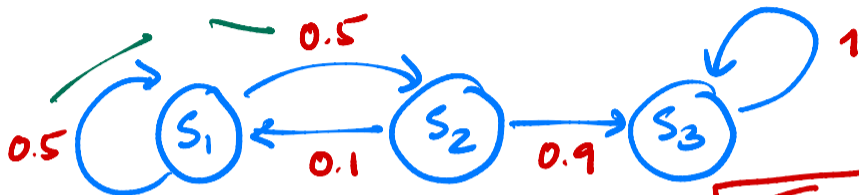
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 $0.5 \times 0.9 = 0.45$
- 0.45 is **likelihood weight** of sample
- Samples s_1, s_2, \dots, s_N with weights w_1, w_2, \dots, w_N

$$\blacksquare P(r \mid c, w) = \frac{\sum_{s_i \text{ has rain}} w_i}{\sum_{1 \leq j \leq N} w_j}$$



Markov chains



Add up to 1

Each row sums to 1

Stochastic Matrix

Transition Matrix

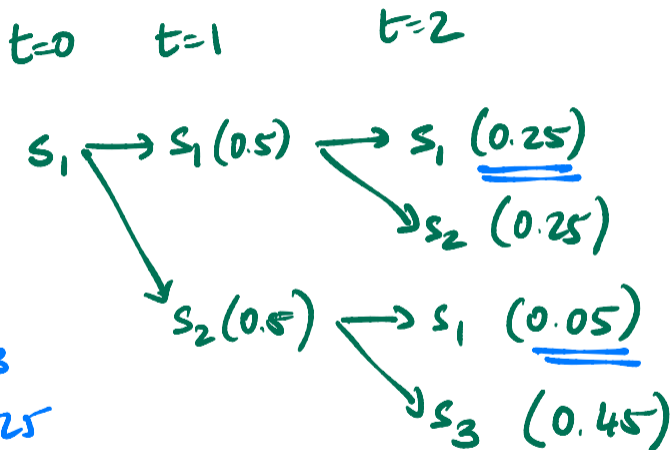
	s_1	s_2	s_3
s_1	0.5	0.5	0
s_2	0.1	0	0.9
s_3	0	0	1

Markov chains

	s_1	s_2	s_3
s_1	0.5	0.5	0
s_2	0.1	0	0.9
s_3	0	0	1

At $t=2$

$$P(s_1) = 0.3$$
$$P(s_2) = 0.25$$
$$P(s_3) = 0.45$$



State Distribution

$$\begin{array}{l}
 s_1 \\
 s_2 \\
 s_3
 \end{array}
 \begin{bmatrix} 1 \\ 0 \\ 0 \end{bmatrix}
 \rightarrow
 \begin{bmatrix} 0.5 \\ 0.5 \\ 0 \end{bmatrix}
 \rightarrow
 \begin{bmatrix} 0.3 \\ 0.25 \\ 0.45 \end{bmatrix}
 \rightarrow \dots$$

What happens in the limit?

In this example, intuitively

$$\begin{bmatrix} 0 \\ 0 \\ 1 \end{bmatrix}$$

Markov chains

	s_1	s_2	s_3
s_1	0.5	0.5	0
s_2	0.1	0	0.9
s_3	0	0	1

M

state transition matrix

π - state distribution

$$\pi_0 = [1 \ 0 \ 0]$$

$$\pi_1 = ?$$

$$\pi_0 \cdot M$$

$$\pi_q [j] = \sum_{k=1}^3 \pi_0 \cdot M_{kj}$$

$$\pi_{k+1} = \pi_k M$$

Stationary distribution

$$\pi^* = \pi^* M$$

M has a stationary distribution if it is ergodic

- irreducible / strongly connected
- aperiodic

Periodic



Aperiodic: $\exists k$ s.t. $\forall s_i, s_j$ there exist paths of length $k, k+1, k+2, \dots$ from s_i to s_j

Claim If M is ergodic, M has a ^{unique} stationary distribution π

$$\pi M = \pi$$

"Compute" π^* by running the Markov chain
long enough from any initial state

Stationary distribution represents fraction of visits to each state in a long enough execution

How fast do we converge to π^* ?

"Mixing" time

Bayesian network variables V_1, V_2, \dots, V_n

Each assignment of values to V_i 's is a sample

Samples x_1, x_2, \dots, x_N - all possible configurations of V_i 's

\downarrow \downarrow \downarrow
 p_1 p_2 p_N

$$\sum p_i = 1$$

Design a Markov chain with N states

$S_i \longleftrightarrow X_i$

\Downarrow

M

\downarrow

π^*

st. $\pi_i^* = P_i$